Investigating the association between educational attainment and allostatic load with risk of cancer mortality among African American women

Cynthia Li¹, Sydney Elizabeth Andrzejak¹, Samantha R. Jones², Brittany Marie Williams³ and Justin Xavier Moore⁴,⁵,⁶*

Abstract

Background African American (AA) women navigate the world with multiple intersecting marginalized identities. Accordingly, AA women have higher cumulative stress burden or allostatic load (AL) compared to other women. Studies suggest that AA women with a college degree or higher have lower AL than AA women with less than a high school diploma. We examined the joint effect of educational attainment and AL status with long-term risk of cancer mortality, and whether education moderated the association between AL and cancer mortality.

Methods We performed a retrospective analysis among 4,677 AA women within the National Health and Nutrition Examination Survey (NHANES) from 1988 to 2010 with follow-up data through December 31, 2019. We fit weighted Cox proportional hazards models to estimate adjusted hazard ratios (aHRs) of cancer death between educational attainment/AL (adjusted for age, income, and smoking status).

Results AA women with less than a high school diploma living with high AL had nearly a 3-fold increased risk (unadjusted HR: 2.98; 95% CI: 1.24–7.15) of cancer death compared to AA college graduates living with low AL. However, after adjusting for age, this effect attenuated (age-adjusted HR: 1.11; 95% CI: 0.45–2.74). AA women with high AL had 2.3-fold increased risk of cancer death (fully adjusted HR: 2.26; 95% CI: 1.10–4.57) when compared to AA with low AL, specifically among women with high school diploma or equivalent and without history of cancer.

Conclusions Our findings suggest that high allostatic load is associated with a higher risk of cancer mortality among AA women with lower educational attainment, while no such association was observed among AA women with higher educational attainment. Thus, educational attainment plays a modifying role in the relationship between allostatic load and the risk of cancer death for AA women. Higher education can bring several benefits, including improved access to medical care and enhanced medical literacy, which in turn may help mitigate the adverse impact of AL and the heightened risk of cancer mortality among AA women.

Keywords Cancer, life-course, Cumulative stress, Psychosocial stress, Race, Disparities, Allostatic load
Introduction

Chronic stress has been linked to worse health outcomes for various diseases, from atherosclerosis and hypertension to depression to hypertension [1, 2]. The body undergoes physiologic changes to compensate for chronic stress. There is an increased production of corticosteroids and catecholamines, an increased inflammatory response, increased levels of oxidative stress, and DNA damage [3, 4]. Dai et al. suggest that this cumulative effect and physiologic changes promote tumorigenesis and cancer development by suppressing immunity and enhancing inflammation [3].

Allostatic load (AL) measures the cumulative burden of chronic stress on physiological systems. Chronic stress can result from major life events, but more importantly, is the summative effect of ordinary, everyday stressors such as poor sleep, a lack of exercise, and a lack of access to healthy food [5]. A cross-sectional study by Allen et al. found that racial discrimination may be an important predictor of cumulative physiologic dysregulation. [6] Experiences of racial discrimination as a stressor have been correlated with worse physical and mental health outcomes [7–9]. The “weathering” hypothesis by Geronimus et al. suggested that systemic stress from cumulative socioeconomic disadvantage and political marginalization contribute greatly to the early and disproportionate health deterioration of African Americans (AA) [10].

Multiple studies have illuminated that AA men and women consistently have the highest allostatic load scores, compared to those of the same age in other racial groups [10–12]. AA women navigate this world with multiple, intersectional marginalized identities, meaning they are subject to multiple forms of layered discrimination by race, sex, class, and other social group categories. AA women consistently have higher allostatic load compared to their AA male and White female counterparts [6, 13]. A longitudinal study by Upchurch et al., observed that AA women reported higher levels of discrimination, perceived stress, and hostility, all of which will increase allostatic load [14]. In a cross-sectional study, Moore et al. found that AA women had the highest age-adjusted mean allostatic load scores at the end of 30 years compared to others in their same age group [15]. Considering that chronic stress and thus allostatic load has been associated with biological changes that promote tumorigenesis, it is worthwhile to understand the correlations between allostatic load in AA women and cancer mortality [3, 16, 17].

Furthermore, Williams et al. found that AA women with lower educational attainment had a higher allostatic load [18]. Therefore, considering educational attainment’s impact on allostatic load is important in understanding the specific effects of chronic stress on health outcomes. AL and chronic stress are directly associated with worse health outcomes in multiple disease states, particularly cancer [3, 16, 17]. The present study explored the joint effect of educational attainment levels and allostatic load on the association with long-term risk of cancer mortality, and the moderating role of educational attainment on the association between allostatic load and cancer mortality.

Methods

Study design and participants

We performed a retrospective cohort analysis using data from the National Health and Nutrition Examination Survey (NHANES), a representative sample of noninstitutionalized US residents linked with the National Center for Health Statistics (NCHS) 2019 National Death Index (NDI) file. The NHANES program oversamples those aged 60 and older, Latinx and non-Hispanic (NH)-Black (henceforth, African American or AA) individuals, and weighted analysis generates generalizable estimates [19]. The weighted sample of NHANES is comparative to non-institutionalized United States (U.S.) population [20]. Using NHANES survey data from years 1988 through 2010 linked with NDI data (follow-up data through December 31, 2019) we examined the association between the intersectionality of educational attainment and allostatic load with risk of cancer mortality. The NHANES survey includes information on sociodemographics, clinical measurements, and health-related questionnaires. NHANES participants with data on biomarkers were used in this analysis. We performed analysis among NHANES participants with data on biomarkers and within a fasting subsample (N=95,359). Patients were excluded if they reported current pregnancy or were less than 18 years of age (N=42,791), were missing AL biomarkers or not linked via NDI (N=11,360). This resulted in a final analytic sample of NHANES participants aged 18 and older, corresponding to a total of 41,218 participants over a 22-year study period, of which 4,677 identified as AA women (Fig. 1).

We completed analyses using domain statements to account for appropriate estimations of covariance-variance structures using specific strata, cluster, and weighting procedures as specified by NHANES methodology. We created a race/ethnicity-sex specific variable based on the intersectionality of race/ethnicity with sex specified at survey, and thus we had an eight level variable containing: [1] AA men, [2] AA women, [3] NH-White men, [4] NH-White women, [5] Hispanic men, [6] Hispanic women, [7] Other/mixed race men, and [8] Other/mixed race women. We also note that NHANES did not collect granular gender identity terms, and thus we are unable to elucidate whether men/women were cis- or trans- identifying adults. Mortality status or vital status
Ethical statement
The Institutional Review Boards considered this study exempt from review because of the use of secondary, publicly available, and de-identified data.

Educational attainment, effect modifier
This study mirrored methods of our investigative team’s prior work, Williams et al. (2022), to determine our variables of interest [18]. We examined educational attainment as an effect modifier on the relationship between which was determined from the NHANES question “What is the highest grade or level of school you completed or highest degree received?”. We then categorized educational attainment into a four-level variable based on participants’ that completed [1] less than an high school (HS) education; [2] high school graduate, general education development test (GED), or equivalent; [3] some college; and [4] college graduate or above [18]. Due to the NHANES data collection, we could not differentiate by specific degree types (e.g., MD, PhD, MSN). Educational attainment was treated as an effect modifier within our education stratified results examining the association...
between allostatic load and long-term risk of cancer death.

**Allostatic load, primary independent variable**

Allostatic load has been defined using varying components, although most incorporate biomarker measures from three different categories, including physiologic functioning, which incorporates cardiovascular, metabolic, and immune systems [21]. While there is no consensus definition, we decided to define AL using the Geronomus et al. (2006) and Moore et al. (2021) taxonomies [10, 15]. AL components included body mass index (BMI), diastolic blood pressure (DBP), mean arterial pressure (MAP), total cholesterol, serum triglycerides, serum albumin, serum creatinine, and C-reactive protein (CRP). We considered sex as a biological variable according to National Institutes of Health guidelines regarding human subjects research [22, 23]. To determine the high-risk thresholds for each AL component, we examined the sex reported at survey-specific cutoffs of each component among the entire study sample with complete biomarker data. High-risk thresholds were determined by either being above the 75th percentile for BMI, CRP, DBP, glycated hemoglobin, SBP, total cholesterol, serum triglycerides, and serum creatinine [24, 25] or below the 25th percentile for serum albumin. Therefore, each NHANES participant was scored as either 1 (high-risk) or 0 (low-risk) based on sex at baseline survey-specific cutoffs for each component. Total AL score was calculated by summing the individual components, ranging from 0 to 9. Participants were further categorized with AL scores greater or equal to 3 as having high AL [21, 26].

**Joint effect of allostatic load and educational attainment**

After categorizing NHANES participants based on the distribution of AL components and their self-reported educational attainment, we created a variable examining the intersection of AL and educational attainment. This variable was categorized into eight levels; [1] college graduate or more living with low AL (n=256), [2] college graduate or more living with high AL (n=258), [3] some college with low AL (n=524), [4] some college with high AL (n=598), [5] HS diploma or equivalent with low AL (n=662), [6] HS diploma or equivalent with high AL (n=849), [7] less than HS with low AL (n=473), and [8] less than HS with high AL (n=1,044).

**Primary outcome of interest, cancer death**

Our primary outcome of interest was the time to cancer death. Deaths attributed to malignant neoplasms (ICD-10 18 019–043) were included as cancer-related deaths. Our primary outcome of interest was time to cancer-related death. Follow-up data for this analysis was available through December 31, 2019, based on NDI-NHANES publicly available linkages. The primary determination of mortality for eligible NHANES participants is based upon matching survey records to the NDI although additional redundant sources are also incorporated, including the Social Security Administration, the Centers for Medicare and Medicaid Services, data collection, NCHS’ follow-up surveys (e.g., NHEFS), and ascertainment of death certificates.

**Other variables of interest**

We included other variables as covariates based on their consideration as potential confounders, or their possible effect on education, cumulative stress, and cancer outcomes based on prior studies. These variables included NHANES baseline survey completion period (e.g., 1988–1991 through 2009–2010), family poverty-to-income ratio (PIR), current smoking status, any self-reported history of cancer, congestive heart failure, and ever heart attack. PIR was calculated as the ratio of total family income to poverty threshold values by NHANES investigators [18]. Participants that reported no income were given a zero value for PIR [18]. PIR values greater than 1 are above the poverty level, and values near 5 are considered very high income, while PIR values less than 1 are considered below the official poverty line [18]. Participants who had smoked at least 100 cigarettes in their lifetime and who were currently smoking during survey administration were categorized as current smokers [18]. We determined self-reported diagnoses by doctor for cancer, congestive heart failure, or heart attack from NHANES questionnaires on whether “…a doctor or other health professional ever told you that you had … (cancer, angina, congestive heart failure (CHF), or heart attack” [18].

**Statistical analysis**

Primary analyses were conducted using NHANES-generated sampling statistical strata, clusters, and weights as designated and described in detail within the NHANES methodology handbook [27]. NHANES only measures biomarkers among a random sample of participants each survey period, and in turn creates subsample weights to account for the probability of being selected into the subsample component and additional non-response bias. We combined NHANES weighting variable accounting for the 18-years of cross-sectional data (1988–1994 and 1999–2010). The mobile examination center (MEC) included physical measurements such as blood pressure, a dental examination, and the collection of blood and urine specimens for laboratory testing. Following analytic guidelines by the National Center for Health Statistics (NCHS) [20, 28, 29], we utilized the NHANES MEC sample weights for NHANES III (years 1988–1994)
and NHANES 1999–2010 and constructed an adjusted weight by modifying the weights to have a common denominator of 18; that is, [1] the six-year weight for 1988–1994 was multiplied by 1/3, [2] the four-year weight for 1999–2002 was multiplied by 1/4.5, and [3] the two-year cycle weights for each subsequent NHANES survey years 2003–2010 were multiplied by 1/9. There are three methodological assumptions to note when combining these years of data: [1] there are no differences in the estimates over the time periods being combined; [2] regarding the interpretation, the estimates are the average over the period; and [3] NHANES III (1988–1994) recruited US non-institutionalized population aged 2 months and older, while continuous (1999 and later) has recruited all ages.

For descriptive statistics we presented categorical variables as weighted row percentages and continuous variables as mean and associated 95% confidence intervals using appropriate SAS survey weighted procedures including PROC SURVEYFREQ, SURVEYMEANS, and SURVEYREG accordingly [30]. Mean survival times were estimated using the product-limit method of the Kaplan-Meier survival estimator. Proportionality assumption was assessed for our primary variable of interest (educational attainment by allostatic load status) by examining the proportion of 1000 simulations that contain a maximum cumulative martingale residual larger than the observed maximum cumulative residuals using the SAS procedure ‘supremum test’. None of our exposure levels had p values that were statistically significant (p value < 0.05), and therefore none of our residuals were larger than expected and we did not reject proportional hazards assumptions. Relative rates of cancer death by groups of educational attainment/allostatic load were estimated by fitting survey-weighted Cox proportional hazards models with time-to-cancer death as the endpoint [16]. Individuals were censored at the time of their event, death, or end of follow-up (December 31, 2019). Models were sequentially adjusted first for age, then with age, poverty to income ratio, and smoking status.

Multiplicative interactions of AL and educational attainment were examined by introducing an interaction term within our model and presenting the corresponding p-value for this association. P-values ≤ 0.05 were considered statistically significant. Additionally, we conducted all the time-to-cancer death event survival analyses by allostatic load status (high versus low allostatic load), stratified by educational attainment. Estimates were presented from our survey-weighted Cox proportional hazard models as hazard ratios (HRs) and associated 95% confidence intervals (CIs). We conducted two sets of sensitivity analyses. We repeated all survival analyses using un-weighted Cox proportional hazard models and thus treating NHANES participants as a simple random sample. In addition, we examined the relationships between the joint effect of education with allostatic load on cancer death risk, and the moderated effect of education on the relationship between allostatic load and cancer death risk while excluding NHANES participants with a history of cancer. All statistical analyses were performed using SAS (version 9.4, SAS Institute, Inc., Cary, North Carolina, USA).

Results
Among 4,677 (an estimated 9,381,049) AA women, the average age of participants was 42.71 years (Standard Error (SE) = 0.33) and the median follow-up time was 15.90 years (Q1-Q3 = 11.48–21.88) (Table 1). AA women with less than high school education attainment living with high allostatic load were on average older (55.5 years, SE = 0.7) than all other groups, while those with high school diploma or equivalent education attainment living with low allostatic load were on average younger (32.8 years, SE = –0.6) than all other groups (p value < 0.01). Participants with college graduate or more educational attainment living with high allostatic load (mean PIR = 3.6, SE = 0.1) and living with low allostatic load (mean PIR = 3.5, SE = 0.1) had much higher income than all other groups (p value < 0.01). The group of participants with highest rate of current smoker status were those with less than high school educational attainment and living with high allostatic load (n = 267, 28.2%). Generally, participants living with high allostatic load (4.9%, 5.0%, 5.1%, 5.3%) were more likely to have a history of cancer compared to those living with low allostatic load (1.9%, 2.8%, 2.6%, 2.5%), among those with less than high school, high school or equivalent, some college, and college graduates or more, respectively (p value < 0.01). Participants with less than high school attainment and high allostatic load had the highest prevalence of congestive heart failure (6.2%) and history of heart attack (5.1%).

**Joint effect of educational attainment with allostatic load**
AA women with an educational attainment of less than high school living with high allostatic load had nearly a 3-fold increased risk (unadjusted HR: 2.98; 95% CI: 1.24–7.15) of dying from cancer compared to college graduates living with low allostatic load (Table 2). However, after adjusting for age, we observed that this association attenuated (age-adjusted HR: 1.11; 95% CI: 0.45–2.74). We observed no other statistically significant associations between educational attainment groups and allostatic load status with risk of cancer death.

**Moderating effect of educational attainment on AL and cancer death risk**
However, when stratified by educational attainment, we observed that participants with less than a high school
Living with Low Allostatic Load

Table 1 Socio-demographic characteristics, personal health, and medical conditions by allostatic load and educational attainment, National Health Examination Survey (NHANES) study period. Among 4,677 NHANES survey participants (an estimated 9,381,049 non-institutionalized African American women) years 1988 through 2010 and follow up through December 31, 2019

<table>
<thead>
<tr>
<th>Time Period</th>
<th>&lt; High School</th>
<th>High School or Equivalent</th>
<th>Some College</th>
<th>College Graduate or more</th>
<th>&lt; High School</th>
<th>High School or Equivalent</th>
<th>Some College</th>
<th>College Graduate or more</th>
</tr>
</thead>
<tbody>
<tr>
<td>1988–1991</td>
<td>245 (21.0)</td>
<td>182 (21.6)</td>
<td>82 (9.9)</td>
<td>40 (4.3)</td>
<td>82 (8.4)</td>
<td>146 (17.2)</td>
<td>88 (12.1)</td>
<td>32 (4.8)</td>
</tr>
<tr>
<td>1991–1994</td>
<td>342 (24.8)</td>
<td>347 (27.6)</td>
<td>143 (11.1)</td>
<td>72 (6.4)</td>
<td>77 (5.2)</td>
<td>171 (13.3)</td>
<td>90 (7.8)</td>
<td>42 (3.6)</td>
</tr>
<tr>
<td>1999–2000</td>
<td>79 (18.5)</td>
<td>39 (10.4)</td>
<td>34 (11.2)</td>
<td>13 (3.5)</td>
<td>68 (16.7)</td>
<td>47 (12.9)</td>
<td>48 (17.0)</td>
<td>26 (8.9)</td>
</tr>
<tr>
<td>2001–2002</td>
<td>63 (13.9)</td>
<td>46 (12.4)</td>
<td>53 (15.7)</td>
<td>21 (5.9)</td>
<td>67 (16.0)</td>
<td>57 (13.2)</td>
<td>49 (15.5)</td>
<td>21 (7.3)</td>
</tr>
<tr>
<td>2003–2004</td>
<td>69 (17.0)</td>
<td>51 (12.8)</td>
<td>51 (15.3)</td>
<td>25 (6.9)</td>
<td>47 (10.6)</td>
<td>68 (12.9)</td>
<td>53 (16.6)</td>
<td>22 (7.7)</td>
</tr>
<tr>
<td>2005–2006</td>
<td>74 (15.0)</td>
<td>59 (12.0)</td>
<td>70 (17.5)</td>
<td>28 (7.2)</td>
<td>49 (9.0)</td>
<td>60 (9.3)</td>
<td>74 (21.1)</td>
<td>34 (8.8)</td>
</tr>
<tr>
<td>2007–2008</td>
<td>94 (15.5)</td>
<td>69 (12.9)</td>
<td>75 (15.4)</td>
<td>33 (7.5)</td>
<td>45 (9.5)</td>
<td>51 (10.7)</td>
<td>63 (15.4)</td>
<td>52 (13.0)</td>
</tr>
<tr>
<td>2009–2010</td>
<td>68 (14.0)</td>
<td>56 (12.4)</td>
<td>90 (20.5)</td>
<td>26 (6.1)</td>
<td>38 (9.5)</td>
<td>62 (15.0)</td>
<td>59 (15.4)</td>
<td>27 (7.1)</td>
</tr>
<tr>
<td>Mean Family PIR</td>
<td>1.5 (0.05)</td>
<td>1.9 (0.06)</td>
<td>2.4 (0.08)</td>
<td>3.6 (0.1)</td>
<td>1.4 (0.07)</td>
<td>1.7 (0.06)</td>
<td>2.3 (0.08)</td>
<td>3.5 (0.1)</td>
</tr>
<tr>
<td>Current smoker status</td>
<td>267 (28.2)</td>
<td>215 (26.8)</td>
<td>113 (17.4)</td>
<td>35 (13.1)</td>
<td>136 (24.6)</td>
<td>145 (25.4)</td>
<td>112 (20.7)</td>
<td>19 (6.4)</td>
</tr>
<tr>
<td>Any cancer history</td>
<td>53 (9.9)</td>
<td>38 (5.0)</td>
<td>35 (5.1)</td>
<td>14 (5.3)</td>
<td>8 (1.9)</td>
<td>14 (2.8)</td>
<td>12 (2.6)</td>
<td>6 (2.5)</td>
</tr>
<tr>
<td>CHF</td>
<td>64 (6.2)</td>
<td>25 (2.9)</td>
<td>20 (3.2)</td>
<td>5 (1.9)</td>
<td>9 (1.7)</td>
<td>3 (0.3)</td>
<td>3 (0.6)</td>
<td>1 (0.2)</td>
</tr>
<tr>
<td>Ever Heart attack</td>
<td>54 (5.1)</td>
<td>27 (2.8)</td>
<td>22 (3.4)</td>
<td>5 (1.8)</td>
<td>8 (1.7)</td>
<td>3 (0.3)</td>
<td>5 (0.9)</td>
<td>1 (0.5)</td>
</tr>
</tbody>
</table>

*High Allostatic load is defined as total Allostatic load score greater than or equal to 3 (presented as column percentages and standard errors).

<Unweighted sample size.

*Presented as unweighted column sample size (weighted percentage) or mean (standard error) for continuous variables.

*Presented as unweighted row sample size (weighted percentage).

Defined as self-reported response to ever being diagnosed by a doctor or health professional of any cancer or malignancy.

degree with high allostatic load had an approximately 3-fold increase in risk of cancer death when compared to those with low allostatic load (unadjusted HR: 3.28; 95% CI: 1.88–5.73). Similarly, among those with a high school diploma or equivalent, there was over a 3.5-fold increase in risk of cancer death in those with high allostatic load compared to those with a low allostatic load (unadjusted HR: 3.61; 95% CI: 1.92–6.76). However, when adjusted for age, the risk of cancer death was attenuated for both the less than high school degree and high school diploma or equivalent groups (<HS age-adjusted HR: 1.41; CI 0.74–2.70; HS diploma or equivalent age-adjusted HR: 1.76, CI: 0.94–3.35). Lastly, among participants with college graduate degree or more, those with high allostatic load had a nearly 2-fold increase in cancer death compared to those with low allostatic load (unadjusted HR: 1.94; 95% CI: 1.39–2.71). This effect, too, was similarly attenuated when adjusted for age (age-adjusted HR: 0.82; 95% CI:
### Table 2
Survey weighted Cox proportional hazard models presented as Hazard Ratios (HR) and 95% Confidence Intervals (CI) for the association between educational attainment/allostatic load and risk of cancer death, among 4,677 (weighted N = 9,381,049) NHANES African American women with 241 (weighted n = 394,768) cancer-related deaths

<table>
<thead>
<tr>
<th>Educational Attainment and Allostatic Load Status</th>
<th>No. &amp; (Weighted %) Cancer Deaths</th>
<th>Mean Survival Months (SE)</th>
<th>Hazard Ratio (HR) and 95% Confidence Interval (CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Unadjusted</td>
<td>Age Adjusted</td>
<td>Fully Adjusted</td>
</tr>
<tr>
<td>College graduate or more with low allostatic load</td>
<td>6 (2.7)</td>
<td>308.6 (1.9)</td>
<td>1.00 (Referent)</td>
</tr>
<tr>
<td>College graduate or more with high allostatic load</td>
<td>12 (44.4)</td>
<td>275.5 (2.0)</td>
<td>1.62 (0.49–5.30)</td>
</tr>
<tr>
<td>Some college with low allostatic load</td>
<td>17 (2.8)</td>
<td>246.2 (1.2)</td>
<td>1.02 (0.32–3.26)</td>
</tr>
<tr>
<td>Some college with high allostatic load</td>
<td>28 (3.4)</td>
<td>235.8 (1.2)</td>
<td>1.32 (0.48–3.66)</td>
</tr>
<tr>
<td>HS diploma or equiv. with low allostatic load</td>
<td>12 (1.7)</td>
<td>260.7 (0.7)</td>
<td>0.55 (0.21–1.45)</td>
</tr>
<tr>
<td>HS diploma or equiv. with high allostatic load</td>
<td>55 (6.1)</td>
<td>251.8 (1.3)</td>
<td>1.97 (0.75–5.17)</td>
</tr>
<tr>
<td>&lt;HS with low allostatic load</td>
<td>15 (2.5)</td>
<td>290.7 (1.7)</td>
<td>0.90 (0.33–2.44)</td>
</tr>
<tr>
<td>&lt;HS with high allostatic load</td>
<td>96 (7.5)</td>
<td>294.8 (2.1)</td>
<td>2.98 (1.24–7.15)</td>
</tr>
</tbody>
</table>

### Sensitivity analyses
We conducted sensitivity analysis treating NHANES participants as simple random sample in an unweighted analysis to examine the joint effect (educational attainment with allostatic load) and moderated effects of educational attainment on the relationship between high allostatic load and risk of cancer death. We observed similar effect measures as main analysis; most noticeably when limited to AA women with less than high school education, those with high allostatic load had 3-fold (unadjusted HR: 3.23; 95% CI: 1.87–5.57) increased risk of cancer death compared to AA women with low AL (Supplemental Table 1). Further, among AA women with high school diploma or equivalent, those with high allostatic load had nearly 4-fold (unadjusted HR: 3.72; 95% CI: 1.99–6.96) increased risk of cancer death when compared to women with low allostatic load.

In an additional sensitivity analysis, we examined the relationships between the joint effect of education with allostatic load on cancer death risk, and the moderated effect of education on the relationship between allostatic load and cancer death risk while excluding NHANES participants with a history of cancer. When stratified among AA women with high school educational attainment or GED equivalent; AA women with high allostatic load had a 4.5-fold increased risk of cancer death (unadjusted HR: 4.54, 95% CI: 2.11–9.74), a 2.3-fold increased risk when accounting for age (age-adjusted HR: 2.27, 95% CI: 1.08–4.74), and 2.3-fold increased risk when fully adjusted (adjusted HR: 2.26, 95% CI: 1.10–4.57) when compared to their counterparts with low allostatic load (Supplemental Table 2). These findings suggest that high allostatic load is strongly associated with risk of cancer death in AA women with lower educational attainment.

### Discussion
In this study, we examined the relationship between educational attainment and allostatic load in AA women and its association with a long-term risk of cancer mortality.
Our findings suggest that high allostatic load, or ‘wear and tear’ because of life course stress, is strongly associated with risk of cancer death in AA women with lower educational attainment but not in AA women with higher educational attainment. AA women with education attainment of less than high school diploma and living with high allostatic load had nearly a 3-fold increased risk of dying from cancer when compared to college graduates living with low allostatic load. However, after adjusting for age, the association between participants having less than high school educational attainment and living with high AL attenuated with the risk of cancer death. When stratified by educational attainment status, we observed that AA women with less than high school and a high school degree or equivalent living with high AL had more than 3-fold increased risk of dying from cancer. However, AA women with some college or college graduates or more with high AL had only a 1–2 increased risk of cancer mortality. We observed that this relationship also reduced after adjusting for confounders.

Within our unadjusted model among AA women, educational attainment modified the effect of AL on cancer mortality. Reductions in AL that may be obtained with higher education, including increased access to healthcare and better understanding of the healthcare system, may partially explain the improved cancer mortality. This means many of the perceived health benefits of a college education do partially result in improved health outcomes for AA women with cancer. Previous studies have found that socioeconomic factors alone are insufficient to fully explain the effect of race on cancer outcomes among AA women, consistent with our findings attenuating with age [16, 17]. Socioeconomic variables in conjunction with cultural beliefs and attitudes, may largely account for sustained disparities in cancer mortality among AA women [31, 32]. Further, increased health literacy for AA women with higher education, differences in tumor phenotype, inherited predispositions, comorbidities, and discrimination and bias experienced by AA women may also account for sustained cancer mortality rates [33, 34]. Moore et al.'s study found that despite living in closer proximity to available healthcare services, increased odds of late-stage diagnosis, no receipt of treatment, and risk of breast cancer death were sustained for NH-AA women living in urban environments compared to rural NH-AA women [35]. One explanation to our findings may be a relative homogeneity in AL in AA women. Our standardization of AL was based on relative data from all races from the NHANES survey. Therefore, it may be possible that differences in high and low AL in AA women were minimal, and therefore, it is difficult to establish a difference in cancer mortality between the two groups.

Previously, Williams et al. determined that AA women with a baccalaureate degree or higher had lower AL [18]. This finding further confirmed existing data suggesting that higher education is a social determinant of health [6, 36, 37]. Though we did not have statistical power within our fully-adjusted statistical models, we observed that when unadjusted and among AA women, educational attainment modified the effect of AL on cancer mortality; as women with high AL compared to women with low AL were at a 3-fold increased risk of cancer death among AA women with less than high school education. This effect was reduced when stratified among AA women with a college education, corresponding to only a 90% increased risk of cancer comparing women with high AL versus those with low AL. Further, these findings show that higher education reduces the effect of high AL, or chronic physiologic stress, on cancer death risk. Further, Moore et al. observed that high AL was associated with an increased risk of overall cancer death [15]. Our results mirrored those of Moore et al., finding that AA women with high AL were more likely to have a history of cancer compared to those living with low AL. Accordingly, future researchers should examine whether the racial weathering associated with living while AA in the U.S. wholly obscures the possibilities of educational attainment mitigating cancer death outcomes. Several studies have shown that AA people generally have the highest cancer mortality burden [16, 17]. With our findings, it does appear that increased education changes the rate of cancer mortality for AA women. The added benefits of higher education, including possible increased income and decreased chronic stress, did lead to decreased cancer mortality. There may also be other factors at play. It is possible that AA women's lived experiences with the healthcare system may play a role in attenuating any benefits of higher education and AL may have on cancer mortality. In addition, a study by Hudson et al. found that the process of upward social mobility may not lend itself to improved health outcomes for AA men and women in the same way that upwardly mobile NH White men and women experience [38]. Meaning, upward mobility may instead be associated with greater health burdens for AA compared to their White counterparts. Indeed, upward mobility may increase experiences with racism and decrease social support. Those with higher educational attainment may attenuate any improvements in cancer mortality associated with higher education in other races due to the stress of negotiating these classed spaces while Black. Future scholars may wish to examine this and other physician-patient factors and their possible role in cancer mortality.

The results from this study should be contextualized by the strengths and limitations of our data. One limitation to our study includes a subsample of non-Hispanic Black women. This subsample resulted in a smaller sample size, which is reflected in the wide confidence intervals.
associated with our hazard ratios. A larger sample size would help improve the precision of our data. Further, allostatic load and baseline exposure variables were attained at a single point in time and not re-assessed. We cannot elucidate life course factors and events between exposure and outcomes. Thus, we were unable to include time-varying measures of education, allostatic load, and other possible life-course factors that may influence the etiology cancer mortality (e.g., diagnosis, prognosis, cancer treatments). Further, because of our use of publicly available NDI-linked NHANES data we were unable to disentangle cancer-specific (e.g., breast, colorectal) mortality. There may be other factors in play between the time of the survey and interview and the time they passed from cancer. Another limitation is in the initial collection of the educational data, wherein there was limited disaggregation for those with at least a baccalaureate degree or higher. Our inability to analyze whether differences at the baccalaureate versus the postgraduate levels forced us to make incomplete inferences about educational attainment, AL, and cancer mortality. However, some strengths of our study include using a nationally representative sample. This allows us to generalize our findings better. Further, this is one of the first studies to look at AL and the risk of cancer death, specifically among AA women.

Conclusion
Previous studies have shown that increased allostatic load is associated with increased risk of cancer death. In our study higher educational attainment modified and reduced the risk of cancer mortality for African American women. However, the role of educational attainment on AL differences in cancer mortality among AA women were significantly attenuated once adjusted for age. Our findings reveal the benefits associated with higher education, such as increased access to medical care and better medical literacy improve but do not fully explain AA women’s risk of cancer mortality. Further research is needed to better understand the factors affecting AA women’s lives that may contribute to higher rates of cancer specific mortality and the potential mediating role of AL.

List of abbreviations

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<tr>
<th>Abbreviation</th>
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<tr>
<td>AA</td>
<td>African American</td>
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<td>AL</td>
<td>Allostatic Load</td>
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<td>NHANES</td>
<td>National Health and Nutrition Examination Survey</td>
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<td>aHRs</td>
<td>Adjusted hazard ratios</td>
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<td>NCHS</td>
<td>National Center for Health Statistics</td>
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<td>NH</td>
<td>Non-Hispanic</td>
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<td>U.S.</td>
<td>United States</td>
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<td>HS</td>
<td>High School</td>
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<td>GED</td>
<td>General Education Development test</td>
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<td>PIR</td>
<td>Poverty to Income Ratio</td>
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<td>CHF</td>
<td>Congestive heart failure</td>
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<td>BMI</td>
<td>Body mass index</td>
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<td>DBP</td>
<td>Diastolic blood pressure</td>
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Supplementary Information

The online version contains supplementary material available at https://doi.org/10.1186/s12905-023-02529-3.

Supplemental Table 1: Unweighted Cox proportional hazard models for the association between educational attainment and allostatic load status with risk of cancer death presented as hazard ratios (HRs) and associated 95% confidence intervals (CIs), among 4,677 non-Hispanic Black women participants with 241 cancer-related deaths.

Supplemental Table 2: Excluding 180 NHANES participants with history of cancer surveyed, weighted Cox proportional hazard models presented as Hazard Ratios (HR) and 95% Confidence Intervals (CI) for the association between educational attainment/allostatic load and risk of cancer death, among 4,484 (weighted N = 8,994,349) NHANES participants with 204 (weighted n = 327,898) cancer-related deaths.

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Authors’ contributions
Cl: Conceptualization, Formal Analysis, Investigation, Data Curation, Writing- Original Draft, Writing- Review & Editing. SA: Conceptualization, Validation, Writing- Review & Editing, Project Administration, Visualization. SJ: Conceptualization, Writing- Original Draft, Writing- Review & Editing. BW: Conceptualization, Writing- Original Draft, Writing- Review & Editing, Visualization. JM: Conceptualization, Formal Analysis, Investigation, Resources, Software, Data Curation, Writing- Original Draft, Writing- Review & Editing, Supervision, Funding Acquisition.

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Data availability
The datasets generated or analyzed during the current study are available in the Centers for Disease Control and Prevention National Health and Nutrition Examination Survey (NHANES) repository, https://wwwn.cdc.gov/nchs/nhanes/.

Declarations

Ethical statement
The Institutional Review Boards considered this study exempt from review because of the use of secondary, publicly available, and de-identified data.

Consent for publication
Not applicable

Competing interests
The authors declare no potential conflicts of interest.

Author details
1Georgia Cancer Center, Augusta University, 1410 Laney Walker Blvd, 30912 Augusta, GA, USA
2Department of Family and Community Medicine, Medical College of Georgia at Augusta University, Augusta, GA, USA
3Department of Education, University of Vermont, Burlington, VT, USA
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